# Sparse High Dimensional Linear Regression: Estimating squared error and a Phase Transition

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### Introduction

The Linear Regression Problem:

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Setup: Let  $\beta^* \in \mathbb{R}^p$ . For some measurement matrix  $X \in \mathbb{R}^{n \times p}$ , and noise vector  $W \in \mathbb{R}^n$ , we observe n noisy linear samples of  $\beta^*$ ,  $Y \in \mathbb{R}^n$ , given by

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(Notation: We call p the number of **features** and n the number of **samples**.)

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#### An immediate answer under full generality: at least p.

**Reason:** Even if W = 0, we have  $Y = X\beta^*$ , a linear system with p unknowns and n equations! To solve it, we need at least p equations, i.e.  $n \ge p$ .

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**Question:** Are we doomed to not use all the features or can we handle such a situation?

(1) Sparsity assumption; we assume  $\beta_i^*$  is zero for all  $i \in \{1, 2, ..., p\}$  except a subset of the indices of cardinality  $k \ll p$ .

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#### Appears a lot

- ▶ in applications; e.g. in signal and image coding [Mallat and Zhang '93].
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• Discrete structure  $\Rightarrow$  easier to analyze.

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- Keeps the challenge of support recovery (a highly nontrivial task)

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Less known in the literature, but

- Discrete structure  $\Rightarrow$  easier to analyze.
- Keeps the challenge of **support recovery** (a highly nontrivial task)
- Best known information theoretic lower bound is much smaller than the best known algorithmic upper bound.

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We assume that

(1)  $X_{i,j}$  is i.i.d. standard normal N(0,1) for all  $i=1,2,\ldots,n$  and  $j=1,2,\ldots,p.$ 

(2) W<sub>i</sub> is i.i.d. normal N(0,  $\sigma^2$ ) for i = 1, 2, ..., n, where  $\sigma^2 = o(k)$ .

(3) X, W are independent.

Classic in literature ([Candes, Tao '06], [Donoho '06], [Wainwright '09])

### The New Model

**Setup:** Let  $\beta^* \in \{0, 1\}^p$  be a **binary** k-sparse vector. For

- $X \in \mathbb{R}^{n \times p}$  consisting of entries i.i.d N(0,1) random variables
- W  $\in \mathbb{R}^n$  consisting of entries i.i.d. N(0,  $\sigma^2$ ) random variables with  $\sigma^2 = o(k)$

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we get n noisy linear samples of  $\beta^*$ ,  $Y \in \mathbb{R}^n$ , given by,

$$\mathsf{Y} := \mathsf{X}\beta^* + \mathsf{W}.$$

**Goal:** Given (Y, X), recover  $\beta^*$  with the minimum number of samples. The recovery should happen with probability tending to 1 as the problem parameters tend to infinity **(w.h.p.)**.

• Upper bounds ([Candes, Tao '06],[Donoho '09],[Wainwright '09]) If

 $n>2k\log p$ 

LASSO and other efficient algorithms recover  $\beta^*$  w.h.p. .

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• Lower bounds ([Wang et al '10])

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• Lower bounds ([Wang et al '10])

If  $n < n^* := \frac{2k}{\log\left(\frac{2k}{\sigma^2} + 1\right)} \log p$ , then there is **no recovery mechanism** of  $\beta^*$  which succeeds w.h.p.

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# The Gap $n^* < n < 2k \log p / Main$ Results

#### The next natural question:

Is it **possible** to recover  $\beta^*$  for n with

 $n^* < n < 2k \log p?$ 

If yes, is there an efficient way to make this recovery?

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If yes, is there an efficient way to make this recovery?

**Main Results:** We answer **yes** to the first question, and conjecture (based on geometrical arguments) that the answer is **no** to the second.

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It has a simple-to-state form: the MLE  $\hat{\beta}$  is the optimal solution of

$$(\Phi_2): \min_{\beta \in \{0,1\}^p, \sum_{i=1}^p \beta_i = k} ||\mathbf{Y} - \mathbf{X}\beta||_2.$$

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### Maximum Likelihood Estimator- "All or Nothing" Theorem

#### Definition

For  $\beta \in \{0, 1\}^p$ , k-sparse we define

 $\mathsf{Overlap}(\beta) := |\mathsf{Support}(\beta^*) \cap \mathsf{Support}(\beta)|.$ 

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#### Theorem ("All or nothing")

$$\begin{aligned} & (Gamarnik, Z. \ 2016) \ Set \ \mathsf{n}^* := \frac{2\mathsf{k}}{\log\left(\frac{2\mathsf{k}}{\sigma^2}+1\right)} \log \mathsf{p} \ and \ let \ \epsilon > 0 \ be \ arbitrary. \\ & \bullet \ If \ \mathsf{n} < (1-\epsilon) \ \mathsf{n}^*, \ then \ w.h.p. \ \frac{1}{\mathsf{k}} \mathsf{Overlap}(\hat{\beta}) \to 0, \ as \ \mathsf{n}, \mathsf{p}, \mathsf{k} \to +\infty. \\ & \bullet \ If \ \mathsf{n} > (1+\epsilon) \ \mathsf{n}^*, \ then \ w.h.p. \ \frac{1}{\mathsf{k}} \mathsf{Overlap}(\hat{\beta}) \to 1, \ as \ \mathsf{n}, \mathsf{p}, \mathsf{k} \to +\infty. \end{aligned}$$

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Comments:

(1) Information **exists** when  $n > (1 + \epsilon)n^*!$ 

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Comments:

- (1) Information **exists** when  $n > (1 + \epsilon)n^*!$
- (2) A **sharp** phase transition!

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- (2) A **sharp** phase transition!
- (3) A challenging application of the second moment method.

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**Question:** Why no efficient algorithm is known when  $n^* < n < 2k\log p$  and many are when  $n > 2k\log p$ ?

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A *usual* picture in the analysis of randoms CSPs. Theory of random CSPs suggests that a usual reason is an **"important change in the geometry of the space of solutions"** between the two regimes.[Achlioptas et al, 2008].

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Usually when such a property holds no efficient algorithm exists and when it ceases, even "local" algorithms work (remember yesterday's talk).

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Various names: shattering property, overlap gap property.

# The Overlap Gap Property (OGP) for Linear Regression

**The OGP** (informally): The set of  $\beta$ 's with "small"  $||Y - X\beta||_2$ "shatters" in two components, one where  $\beta$  have low overlap with the ground truth  $\beta^*$  and one where they have high overlap with  $\beta^*$ .



#### Figure: The OGP around Y

# The Overlap Gap Property for Linear Regression-definition

For r > 0, set  $S_r := \{\beta \in \{0,1\}^p : ||\beta||_0 = k, n^{-\frac{1}{2}} ||Y - X\beta||_2 < r\}.$ 

#### Definition (The Overlap Gap Property)

Let r > 0 and  $0 < \zeta_1 < \zeta_2 < 1$ . We say that the high-dimensional linear regression problem defined by  $(X, W, \beta^*)$  satisfies the Overlap Gap Property with parameters  $(r, \zeta_1, \zeta_2)$  if the following holds.

(a) For every 
$$\beta \in S_r$$
,

$$\frac{1}{\mathsf{k}}\mathsf{Overlap}\left(\beta\right) < \zeta_1 \text{ or } \frac{1}{\mathsf{k}}\mathsf{Overlap}\left(\beta\right) > \zeta_2.$$

(b) Both the sets

$$\mathsf{S}_{\mathsf{r}} \cap \{\beta : \frac{1}{\mathsf{k}}\mathsf{Overlap}\,(\beta) < \zeta_1\} \text{ and } \mathsf{S}_{\mathsf{r}} \cap \{\beta : \frac{1}{\mathsf{k}}\mathsf{Overlap}\,(\beta) > \zeta_2\}$$

are non-empty.

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# The Overlap Gap Property- The result

#### Theorem

There exists C > c > 0 such that,

- If  $n^* < n < ck \log p$  then w.h.p. OGP holds for some  $r = r_k$  and  $0 < \zeta_1 < \zeta_2 < 1.$
- If  $n > Ck \log p$  then w.h.p. OGP does **not** hold for any choice of  $r = r_k$  and  $0 < \zeta_1 < \zeta_2 < 1.(post-COLT)$

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An easy **corollary**: if  $n < ck \log p$  then any "local-greedy" algorithm will fail w.h.p.

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An easy **corollary**: if  $n < ck \log p$  then any "local-greedy" algorithm will fail w.h.p.

Also, if  $n > Ck \log p$  then the simplest "local-greedy" works!(post-COLT)

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(1) We show that when  $n > (1 + \epsilon)n^*$  for some  $\epsilon > 0$ , information exists to recover  $\beta^*$ .

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- (2) The performance of the optimal estimator M.L.E. **changes suddenly** w.h.p. when the number of samples crosses the value n\*.

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- (1) We show that when  $n > (1 + \epsilon)n^*$  for some  $\epsilon > 0$ , information exists to recover  $\beta^*$ .
- (2) The performance of the optimal estimator M.L.E. **changes suddenly** w.h.p. when the number of samples crosses the value n<sup>\*</sup>.
- (3) We conjecture that the regime n\* < n < 2k log p is algorithmically hard and we prove a geometrical phase transition to provide support for it.

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- Can it be proven that assuming n < (1 − ε)n\*, there is no information to recover any fraction of the support of β\*?
- Can we prove/provide more support that n\* < n < 2k log p is algorithmically hard? For example, can we find a reduction from the planted clique like in sparse PCA [Berthet, Rigollet '13]?

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# Thank you!!

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,  $\sum_{i=1}^p \beta_i = k$  ( $||Y - X\beta||_2$ ).</sub>

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.</sub>

• For any  $\ell \in \{0,1,\ldots,k\}$  set

$$\mathsf{T}_{\ell} = \{\beta \in \{0,1\}^{\mathsf{p}} | \sum_{i=1}^{\mathsf{p}} \beta_i = \mathsf{k}, \mathsf{Overlap}(\beta) = \ell\}.$$

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• Set  $d_{\ell} = \min_{\beta \in \mathsf{T}_{\ell}} (||Y - X\beta||_2)$ . Then  $d = \min_{\ell=0,1,\dots,k} d_{\ell}$ .

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• We show that w.h.p. for all  $\ell = 0, 1, \dots, k$ ,

$$\mathsf{d}_\ell \sim \sqrt{2\mathsf{k}(1 - \frac{\ell}{\mathsf{k}}) + \sigma^2} \exp\left(-\frac{\mathsf{k}(1 - \frac{\ell}{\mathsf{k}})\log\mathsf{p}}{\mathsf{n}}\right)$$

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• So w.h.p.

$$\mathsf{d} \sim \min_{\ell=0,1,\ldots,k} \mathsf{f}\left(1 - \frac{\ell}{k}\right) \sim \min_{\alpha \in [0,1]} \mathsf{f}(\alpha) \,.$$

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$$\mathsf{n}^* := \frac{2\mathsf{k}}{\log\left(\frac{2\mathsf{k}}{\sigma^2} + 1\right)}\log\mathsf{p}.$$

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So the optimization problem changes behavior exactly at

$$\mathsf{n}^* := \frac{2\mathsf{k}}{\log\left(\frac{2\mathsf{k}}{\sigma^2} + 1\right)}\log\mathsf{p}.$$

Therefore n > n\* iff f is minimized at 1 iff d<sub>ℓ</sub> being minimized at 0, which happens iff the optimal vector has full common support with β\*.

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Two pictures behind the phase transition  $(p = 10^9, k = 10, \sigma^2 = 1, n^* = 136);$ 



**Comment:**  $\alpha := 1 - \frac{\ell}{k}$ , so  $\alpha = 1$  means no recovery and  $\alpha = 0$  full recovery.